

Clay mathematical institute, Oxford

Random polymers and algebraic combinatorics

What an exactly solvable (log-gamma) model tells us
on polymer localization ?

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Joint work with [Vu-Lan Nguyen](#).

Purpose

Directed polymer model \equiv random walk in a random potential.

Weight of a path = exponent of the sum of the potential it has met on its way.

Competition:

$$\underbrace{\text{entropy of paths}}_{\text{diffusive behavior}} \longleftrightarrow \underbrace{\text{disorder strength}}_{\text{inhomogeneities of potential}}$$

Path is attracted by "favorable spots" with large potential values.

Constant potential \implies path spreads smoothly over distances $O(\sqrt{\text{length}})$.

But if **potential has large variability**,

- ☞ path **localizes** on a few corridors with width of order of unity.
- ☞ spreads at "abnormally" large distance (**superdiffusivity**)

Recently, explicitly solvable models of planar polymers (i.e. $(1 + 1)$ -dim.) have been discovered and used to study the second effect typical of the KPZ universality class.

Here, we take the **Log-Gamma polymer** to get sharp results for **localization**.

Outline

- 1 Polymers models
- 2 What is known on localization
- 3 Log-gamma polymer with boundaries
- 4 Main result
- 5 P2P

The Model

Environment: $(\omega(x), x \in \mathbb{Z}^2)$ real, independent identically distributed r.v.'s

Polymer path: over up-right paths $\mathbf{x} = (x_t; 0 \leq t \leq n)$, $x_t - x_{t-1} = \mathbf{e}_1$ or \mathbf{e}_2 .

Point-to-line polymer measure of a path of length n is

$$Q_n^\omega(\mathbf{x}) = \frac{1}{Z_n^\omega} \exp \left\{ \sum_{t=1}^n \omega(x_t) \right\}$$

with $Z_n = Z_n^\omega$ the sum over all up-right paths \mathbf{x} starting at $x_0 = (0, 0)$.

For a typical realization ω of the environment, what is the behavior of the polymer of large size $n \rightarrow \infty$? (asymptotics of the quenched law)

The KPZ universality class

Recent efforts on planar polymer models to understand physics predictions:
KPZ universality class (Kardar, Parisi and Zhang 1986), see Corwin 2012's recent survey.

Non-gaussian scaling limits and statistics, characterized by a few exponents.

KPZ class covers large random matrices, last passage percolation, interacting particle systems (TASEP, ...).

Biblio starting from Johansson (1998 →)

A few explicitly solvable polymer models:

- ⇒ KPZ equation (Quastel 2010, Hairer 2013),
- ⇒ Brownian queues O'Connell-Yor 2003,
- ⇒ log-gamma polymer Seppäläinen 2012.

Log-gamma polymer

Change notations:

$$Y_{i,j} = e^{\omega(i,j)}$$

so that $Z_n = \sum_{\mathbf{x}} \prod_{t=1}^n Y_{x_t}$.

Def.: **Log-gamma polymer** with parameter $\mu > 0$ (abbrev. $\text{Gamma}(\mu)$)

$$Y_{i,j}^{-1} \sim \Gamma(\mu)^{-1} x^{\mu-1} e^{-x} \quad , \quad x > 0.$$

Seppäläinen 2012

Log-gamma polymer is in KPZ class

Seppäläinen 2012 discovered the stationarity property of this model, which makes it explicitly solvable:

- Seppäläinen obtains the value of the free energy

$$n^{-1} \ln Z_n \xrightarrow{n \rightarrow \infty} -\Psi_0(\mu/2), \quad \Psi_0 = \Gamma'/\Gamma$$

and proves that the volume and wandering exponents for fluctuations are

$$\chi = 1/3, \quad \xi = 2/3.$$

- Large deviations of the partition function Georgiou-Seppäläinen 2013
- GUE Tracy-Widom fluctuations for Z_n at scale $n^{1/3}$
Borodin-Corwin-Remenik 2013: for small μ ,

$$n^{-1/3} (\ln Z_n + n\Psi_0(\mu/2)) \xrightarrow{\text{law}} F_{GUE}$$

the GUE Tracy-Widom distribution.

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- Explicit formula for the Laplace transform of the partition function at finite size Corwin-O'Connell-Seppäläinen-Zygouras 2014; integral formula which can be turned into a Fredholm determinant.

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Localization

Can be measured by probability of the "favourite endpoint"
(= largest probability among endpoints)

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In large generality it is proved that the polymer is localized: Carmona-Hu 2002, Comets-Shiga-Yoshida 2003 proved, by (semi)martingale methods, that

$$\lim_n n^{-1} \ln Z_n < \lim_n n^{-1} \ln \mathbb{E} Z_n \iff \liminf_n n^{-1} \sum_{k=1}^n \mathcal{I}_k \geq c_0 > 0.$$

The assumption holds for planar models (Comets-Vargas 2006, Lacoïn 2010). Then,

$$\limsup_{n \rightarrow \infty} \mathcal{I}_n \geq c_0$$

a **sharp** contrast with $O(n^{-d/2})$ with $d = 1$ when $\omega = \text{Cst}$.

General techniques

- ✘ Doob's decomposition relies on (conditional) first moment, which can be estimated by the replica [overlap](#)

$$\frac{1}{n} \sum_{t=1}^n \mathbf{1}_{\{x_t = \tilde{x}_t\}}, \quad x, \tilde{x} \sim Q_{n-1}^\omega$$

independent except they share the same environment. The overlap can be compared to \mathcal{I}_n .

- ✘ In absence of moments, an alternative approach was developed by Vargas 2007 leading directly to \mathcal{I}_n .

Our aim here

Our aim is to find **sharp** asymptotics for localization in the log-gamma model.
Natural questions are:

- ✘ Where does the mass of the polymer measure stem from ?
- ✘ What are the regions which contribute the most ?
- ✘ Are they many ? How far are they ? Are wide are they ?

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✍ In some models, the above questions can be answered in the limit

Temperature $\rightarrow 0$:

Complete localization when the overlap converges to 1.
C-Cranston 2013, C-Yoshida 2013.

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✎ All previous methods are indirect.
In the log-gamma model we have explicit computations.

Connections with Parabolic Anderson model and KPZ equation

The **point-to-point partition function**

$$Z_{m,n} = \sum_{\mathbf{x}: \mathbf{0} \mapsto (m,n)} \prod_{t=1}^{m+n} Y_{x_t},$$

is a solution of the parabolic Anderson equation

$$Z_{m,n} = e^{\omega(m,n)} [Z_{m-1,n} + Z_{m,n-1}]$$

or, equivalently,

$$Z_{m,n} - Z_{m-1,n-1} = [2 - e^{-\omega(m,n)}] Z_{m,n} + (Z_{m-1,n} + Z_{m,n-1} - Z_{m,n} - Z_{m-1,n-1}),$$

prod. Z with a noise

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a discrete version of Stochastic Heat Equation ($t = n + m, r = n - m$, $\mathcal{Z}(t, r) = Z_{n,m}$, and $\eta(t, r) = [\cdot]$ in the above)

$$\frac{\partial}{\partial t} \mathcal{Z}(t, r) = \eta(t, r) \mathcal{Z}(t, r) + \frac{\partial^2}{\partial r^2} \mathcal{Z}(t, r)$$

The Hopf-Cole transform, $\mathfrak{h}(t, r) = \log \mathcal{Z}(t, r)$ solves KPZ equation

$$\frac{\partial}{\partial t} \mathfrak{h}(t, r) = \frac{\partial^2}{\partial r^2} \mathfrak{h}(t, r) + \left(\frac{\partial}{\partial r} \mathfrak{h}(t, r) \right)^2 + \eta(t, r)$$

What we will prove:

When observed on a line $m + n = Cst.$:

- ✘ The population concentrates around the highest peak and spreads at distance $O(1)$ around it.
- ✘ The second high peak does not contribute significantly.
- ✘ In large time, the population density converges (in distribution), without any scaling.

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Boundary Conditions

Assign distinct weight distributions on the boundaries and in the bulk:

$$U_{i,0} = Y_{i,0} \text{ and } V_{0,j} = Y_{0,j} \text{ for } i, j \in \mathbb{N} := \{1, 2, \dots\}.$$

Model b.c.(\(\theta\)): For $\theta \in (0, \mu)$, denote by b.c.(\(\theta\)) the model with

$$\begin{cases} \{U_{i,0}, V_{0,j}, Y_{i,j} : i, j \in \mathbb{N}\} \text{ are independent with distributions} \\ U_{i,0}^{-1} \sim \text{Gamma}(\theta), \quad V_{0,j}^{-1} \sim \text{Gamma}(\mu - \theta), \quad Y_{i,j}^{-1} \sim \text{Gamma}(\mu). \end{cases}$$

Recall the point-to-point partition function

$$Z_{m,n} = \sum_{\mathbf{x}: \mathbf{0} \rightarrow (m,n)} \prod_{t=1}^{m+n} Y_{x_t},$$

and define new weights on horizontal or vertical edges

$$U_{m,n} = \frac{Z_{m,n}}{Z_{m-1,n}} \text{ and } V_{m,n} = \frac{Z_{m,n}}{Z_{m,n-1}}.$$

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Facts (Seppäläinen 2012): Along any down-right path the variables U, V 's are **mutually independent** with marginal distributions

$$U^{-1} \sim \text{Gamma}(\theta), \quad V^{-1} \sim \text{Gamma}(\mu - \theta) \quad .$$

Burke property for pedestrians

$$\begin{aligned}
 U_{1,1}^{-1} &= \frac{Z_{0,1}}{Z_{1,1}} \\
 &= \frac{V_{0,1}}{(V_{0,1} + U_{1,0})Y_{1,1}} \\
 &= \frac{U_{1,0}^{-1}}{U_{1,0}^{-1} + V_{0,1}^{-1}} \times Y_{1,1}^{-1}, \\
 V_{1,1}^{-1} &= \frac{V_{0,1}^{-1}}{U_{1,0}^{-1} + V_{0,1}^{-1}} \times Y_{1,1}^{-1}
 \end{aligned}$$

Recall standard prop.: If u, v are independent $\text{Gamma}(\theta, \text{resp. } \mu - \theta)$, then

$$\left(\frac{u}{u+v}, \frac{v}{u+v} \right) \sim \text{Beta}(\theta, \mu - \theta) \text{ independent of } u+v \sim \text{Gamma}(\mu)$$

And conversely.

Property \implies above Fact.

Random walk representation for ratios of partition functions

Considering the down-right path along the vertices $x : x \cdot (\mathbf{e}_1 + \mathbf{e}_2) = n$, we deduce the representation

$$\frac{Z_{k,n-k}}{Z_{0,n}} = \exp\left(-\sum_{i=0}^k X_i^n\right).$$

with a collection $X_k^n = -\log\left(\frac{U_{k,n-k}}{V_{k-1,n-k+1}}\right)$ of i.i.d.r.v.'s. The endpoint distribution is

$$Q_n^\omega \{x_n = (k, n-k)\} = \frac{Z_{k,n-k}}{\sum_{i=0}^n Z_{i,n-i}} = \frac{1}{\sum_{i=0}^n \exp(-(S_i^n - S_k^n))}$$

with $S_k^n = \sum_{i=1}^k X_i^n$ a **random walk**; It is centered iff $\theta = \mu/2$.

The favorite endpoint is

$$l_n^n = \arg \min \{S_k^n; 0 \leq k \leq n\}.$$

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main result

For every n , consider the end-point distribution centered around favorite endpoint,

$$\hat{\xi}_k^{(n)} = Q_n^\omega \{x_n = (I_n^n + k, n - I_n^n - k)\}, \quad k \in \mathbb{Z}.$$

Thus, $\hat{\xi}^{(n)} = (\hat{\xi}_k^{(n)}; k \in \mathbb{Z})$ is a random element of $\mathcal{M}_1(\mathbb{Z})$.

Theorem

For the model b.c.(\(\theta\)) with $\theta \in (0, \mu)$, we have convergence in law

$$\hat{\xi}^{(n)} \xrightarrow{\mathcal{L}} \xi \quad \text{in the space } (\mathcal{M}_1, \|\cdot\|_{TV}),$$

where $\|\mu - \nu\|_{TV} = \sum_k |\mu(k) - \nu(k)|$ is the total variation distance.

Consequences of main result: answers to our questions

A few consequences:

⇒ **Mass of favourite endpoint** converges

$$\mathcal{I}_n \xrightarrow{\mathcal{L}} \max\{(\xi(k) + \xi(k + 1))/2; k \in \mathbb{Z}\} > 0.$$

Consequences of main result: answers to our questions

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☞ **Tightness of the endpoint**: Letting $\vec{l}_n = (l_n, n-l_n)$,

$$\lim_{K \rightarrow \infty} \limsup_{n \rightarrow \infty} Q_n^\omega[|x_n - \vec{l}_n| \geq K] = 0 \quad \text{in probability}$$

Cf. uniqueness of geodesics, Newman 1995, . . . , Georgiou, Rassoul-Agha, Seppalainen 2015

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☞ **Scaling limit of endpoint:** By Donsker's invariance principle, when $\theta = \mu/2$, the RW S_k can be approximated by a Brownian motion W :

$$\frac{l_n}{n} \xrightarrow{\mathcal{L}} \arg \min_{t \in [0,1]} W_t,$$

the arcsine law.

(And so does $\frac{x_n}{n}$ by previous point.)

☞ When $\theta > \mu/2$, the limit is 0. In fact, if $\theta > \mu/2$, l_n converges in law to a finite limit. And symmetrically if $\theta < \mu/2$.

Consequences of main result, continued

Large deviations of the polymer endpoint:

☞ For $\theta = \mu/2$,

$$Q_n^\omega \{x_n = ([ns], n - [ns])\} \simeq e^{-\sqrt{n}[W(s) - \min_{[0,1]} W]},$$

☞ ... whereas for $\theta > \mu/2$,

$$Q_n^\omega \{x_n = ([ns], n - [ns])\} \simeq e^{-ns|\Psi_0(\theta) - \Psi_0(\mu - \theta)|}.$$

Observe:

- Change of speed in the LDP from equilibrium to non-equilibrium.
- Rate function is random in the first case (it depends on the environment), and deterministic in the second one. (*Differs from Barraquand's talk*)

b.c. are crucial

Observe that these results disagree with KPZ theory, namely on where localization takes place; but we believe that "the trapping at the minimum of a RW" is essentially the correct mechanism.

The disagreement come from the boundary conditions.

sketch of proof of main result for $\theta = \mu/2$

Recall that

$$Q_n^\omega \{x_n = (k, n - k)\} = \frac{1}{\sum_{i=0}^n \exp(-(S_i^n - S_k^n))}$$

Since we are only interested in the law of $Q_n^\omega \{x_n = (k, n - k)\}$, we drop the superscript n in S_n^n, X_i^n, I_n^n , etc. . .

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- **Splitting** a random walk S **at its minimum** is a well studied: Bertoin 1991-94, Doney 1989-94, in the mean-zero case and the drifted case.

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- **Splitting** a random walk S **at its minimum** is a well studied: Bertoin 1991-94, Doney 1989-94, in the mean-zero case and the drifted case.
- For $\theta = \mu/2$ (otherwise, quite different and simpler.): The process converges to 2 independent pieces on \mathbb{Z}^+ and \mathbb{Z}^- , glued at 0,
 - ✂ (i) the first one being the random walk conditioned to stay non negative forever, say S^\uparrow ;
 - ✂ (ii) the other one being the time-reverse of the random walk with jumps $-X$ conditioned to stay strictly positive forever, say S^\downarrow .

For fixed K ,

$$\begin{aligned} (S_{l_n+k} - S_{l_n})_{1 \leq k \leq K} &\xrightarrow{\mathcal{L}} (S_k^\uparrow)_{1 \leq k \leq K}, \\ (S_{l_n+k} - S_{l_n})_{-1 \geq k \geq -K} &\xrightarrow{\mathcal{L}} (S_k^\downarrow)_{1 \leq k \leq K}. \end{aligned}$$

sketch of proof of main result for $\theta = \mu/2$

Since we condition by a null event, (S_k^\uparrow) and (S_k^\downarrow) are taboo processes and a correct definition is via Doob's h -transform.

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Now, to control the full (unbounded) sum, we use a consequence of a result by Ritter 1981:

$$\lim_{\delta \rightarrow 0} \mathbb{P}[S_{k+l_n} - S_{l_n} > \delta k^{1/2-\epsilon} \text{ for all } k \leq n - l_n] = 1 .$$

With the preceding, we get:

$$\frac{1}{\sum_{k=0}^n \exp(-(S_k - S_{l_n}))} \xrightarrow{\mathcal{L}} \frac{1}{\sum_{k=0}^{\infty} \exp(-S_k^\uparrow) + \sum_{k=1}^{\infty} \exp(-S_k^\downarrow)} ,$$

which writes also

$$Q_n^\omega \{x_n = (l_n^n, n - l_n^n)\} = \hat{\xi}_0^n \xrightarrow{\mathcal{L}} \xi_0 .$$

sketch of proof of main result for $\theta = \mu/2$

For general values of k , the same arguments lead to the expression of the limit of $\xi_{l_n+k}^n = \hat{\xi}_k^n$, explicitly:

$$\xi_k = \begin{cases} \frac{\exp(-S_k^\uparrow)}{1 + \sum_{i=1}^{\infty} \exp(-S_i^\uparrow) + \sum_{i=1}^{\infty} \exp(-S_i^\downarrow)}, & \text{if } k \geq 0 \\ \frac{\exp(-S_k^\downarrow)}{1 + \sum_{i=1}^{\infty} \exp(-S_i^\uparrow) + \sum_{i=1}^{\infty} \exp(-S_i^\downarrow)}, & \text{if } k < 0 \end{cases}$$



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Point to point polymer

Fix $\mu > 0$, $(p, q) \in (\mathbb{Z}_+^*)^2$ and for $N \in \mathbb{N}$, let R_N be the rectangle with diagonal $(0, 0)$, (pN, qN) .

Model P2P-b.c.(\(\theta\)): Assume

$$\left\{ \begin{array}{l} Y_{i,j} : (i, j) \in R_N \setminus \{\mathbf{0}, (pN, qN)\} \text{ are independent with} \\ Y_{i,j}^{-1} \sim \text{Gamma}(\theta, 1) \text{ for } j \in \{0, qN\}, \\ Y_{i,j}^{-1} \sim \text{Gamma}(\mu - \theta, 1) \text{ for } i \in \{0, pN\}, \\ Y_{i,j}^{-1} \sim \text{Gamma}(\mu, 1) \text{ for } 0 < i < pN \text{ and } 0 < j < qN. \end{array} \right.$$

Point-to-point polymer measure is the probability measure

$$Q_{pN, qN}^\omega(\mathbf{x}) = \frac{1}{Z_{pN, qN}^\omega} \exp \left\{ \sum_{t=1}^n \omega(x_t) \right\}.$$

For a path \mathbf{x} denote by t^- the "time it crosses the second diagonal". The transverse coordinate of the crossing point can be described by

$$F(\mathbf{x}) = (x_{t^-} + x_{t^-+1}) \cdot (q\mathbf{e}_1 - p\mathbf{e}_2).$$

Point to point polymer

Theorem

For any $\theta \in (0, \mu)$, there exist a random integer m_N depending on ω and a random probability measure $\hat{\xi}$ on \mathbb{Z} such that, as $N \rightarrow \infty$,

$$\left(Q_{pN, qN}^\omega(F(\mathbf{x}) = m_N + k); k \in \mathbb{Z} \right) \xrightarrow{\mathcal{L}} \hat{\xi},$$

in the space $(\mathcal{M}_1, \|\cdot\|_{TV})$.

- ✂ Recall that middle-point localization for the point-to-point measure is not covered by the usual semi-martingale approach to localization, and this result is totally new.

Time is out

Perspectives:

- ✂ Model without boundaries
- ✂ Path localization
- ✂ Other solvable models

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Merci de votre attention !